

AMS
206
31 Jul
19

today: Fisher
next time: Bayes

$$(Y_i | \theta \in \mathcal{B}) \stackrel{IID}{\sim} N(\mu, \sigma^2) \\ (i=1, \dots, n)$$

$$\theta = (\mu, \sigma^2) \in \mathbb{R}^+$$

CS2 Fisher's algorithm for frequentist inference

① (same as Neyman) $\mathbb{C} \rightarrow p(y_i | \theta \in \mathcal{B})$

$$(Y_i | \theta \in \mathcal{B}) \stackrel{IID}{\sim} p(y_i | \theta \in \mathcal{B}) \\ (i=1, \dots, n)$$

ex. (CS2) $(Y_i | \theta \in \mathcal{B}) \stackrel{IID}{\sim}$ Bernoulli(θ) ^{marginal dist sampling}
($i=1, \dots, n$)

② use ① to write down the joint sampling dist.

for $(Y_1, \dots, Y_n) = \mathbf{Y}$ CS2 Y_i discrete

and
 $P(Y_1 = y_1, Y_2 = y_2, \dots, Y_n = y_n | \theta \in \mathcal{B})$

$$P(\mathcal{I}_i = \gamma_i | \theta \mathcal{B}) = \begin{cases} \theta & \text{if } \gamma_i = 1 \\ 1 - \theta & \gamma_i = 0 \end{cases} \quad (2)$$

Bernoulli

$$= \theta^{\gamma_i} (1 - \theta)^{1 - \gamma_i}$$

$$= p(\gamma_i | \theta \mathcal{B})$$

$$P(\mathcal{I}_1 = \gamma_1, \dots, \mathcal{I}_n = \gamma_n | \theta \mathcal{B})$$

$$\stackrel{\text{IID}}{=} \prod_{i=1}^n p(\gamma_i | \theta \mathcal{B})$$

$$\prod_{i=1}^n p(\gamma_i | \theta \mathcal{B}) \stackrel{\text{IID}}{=} \prod_{i=1}^n \theta^{\gamma_i} (1 - \theta)^{1 - \gamma_i}$$

$$= [\theta^{\gamma_1} (1 - \theta)^{1 - \gamma_1}] [\theta^{\gamma_2} (1 - \theta)^{1 - \gamma_2}] \dots [\theta^{\gamma_n} (1 - \theta)^{1 - \gamma_n}]$$

$$= \theta^{\sum_{i=1}^n \gamma_i} (1 - \theta)^{n - \sum_{i=1}^n \gamma_i} \quad \left(s = \sum_{i=1}^n \gamma_i \right)$$

$$P(\mathcal{I}_1 = 1, \dots, \mathcal{I}_n = \gamma_n | \theta \mathcal{B}) = \theta^s (1 - \theta)^{n - s}$$

$$y = (\gamma_1, \dots, \gamma_n)$$

(replace)

③ Define the likelihood function:
 f_n of θ for fixed y (any $n > 0$)

f_n of y for fixed θ

$$l(\theta | y \mathcal{B}) \stackrel{\Delta}{=} c \cdot P(\mathcal{I}_1 = \gamma_1, \dots, \mathcal{I}_n = \gamma_n | \theta \mathcal{B})$$

Ex. 52) $P(I_1 = y_1, \dots, I_n = y_n | \theta = \theta) = \theta^s (1-\theta)^{n-s}$ (3)

$$l(\theta | y, \mathcal{B}) \triangleq c \theta^s (1-\theta)^{n-s}$$

$s = \sum_{i=1}^n y_i$

data-compression

dimensionality reduction from

$y = (y_1, \dots, y_n)$ (n -dimensional)

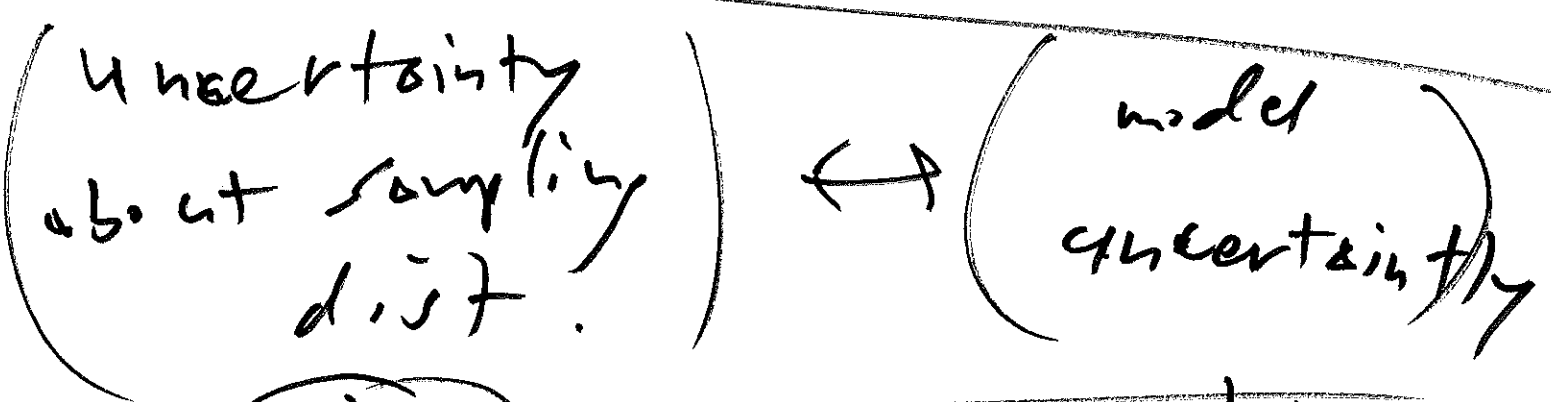
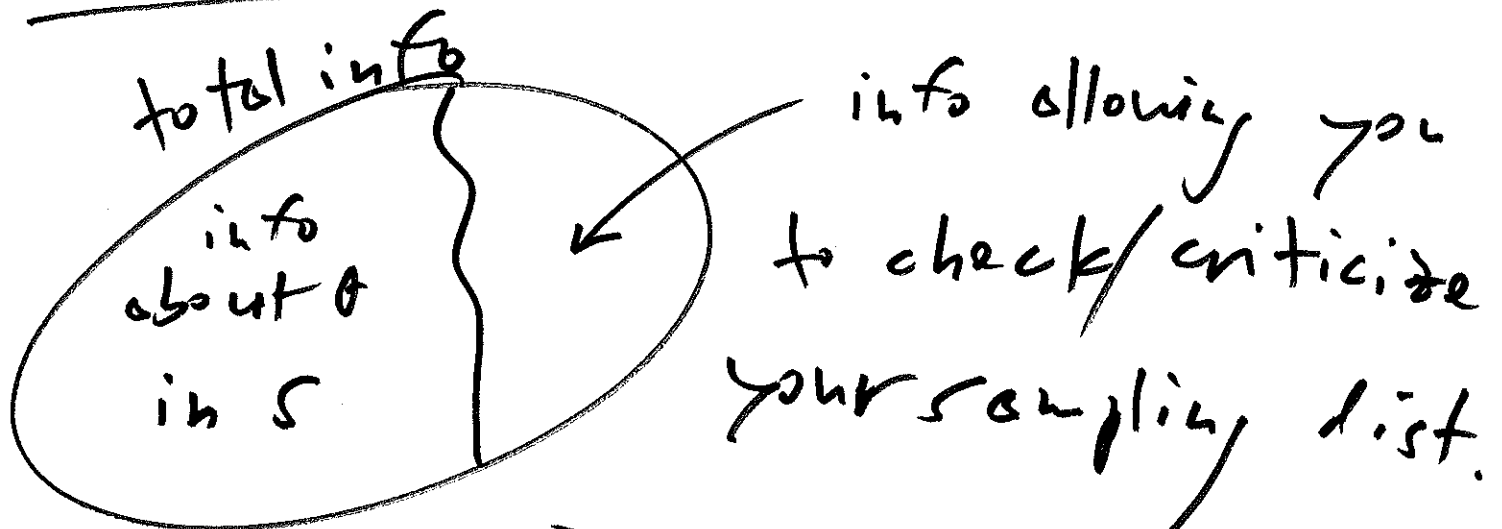
to s (1 -dimensional) (1925)

terabyte 10^{12} bytes
 850
 1,000,000,000,000
 n s

Def. if $l(\theta | y, \mathcal{B})$ depends on y only through g (lower-dimensional)

g (function / summary of y) s ,
 s is said to be a sufficient statistic for θ in this sampling model

Fisher: "when suff. stat^s exists throw y away & keep s " (4) wrong



Fisher horvserhit

$P = (D, C) \xrightarrow{\text{unique}} \underline{p(y, \theta, B)}$

(4) find θ to maximize $l(\theta | y, B)$:

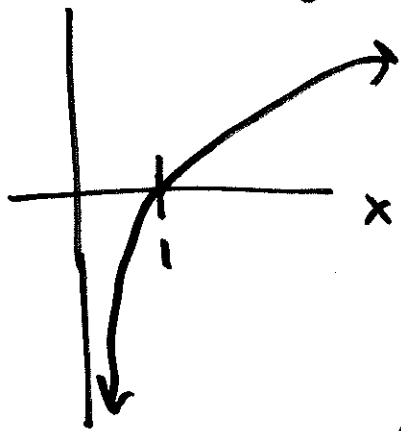
result is $\hat{\theta}_{MLE}$ maximum likelihood estimate

$$\hat{\theta}_{MLE} = \arg \max_{\theta} l(\theta | y, B)$$

④ maximize by likelihood f'_4 : ⑤

$$\hat{\theta}_{MLE} = \arg \max_{\theta} \ell(\theta | \gamma, \mathcal{B})$$

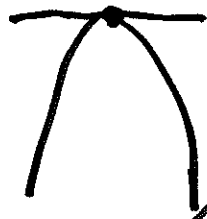
$f(x) = \ln x$
is
monotonic
increasing



Def.

$$\ell(\theta | \gamma, \mathcal{B}) =$$

$$\log \mathcal{L}(\theta | \gamma, \mathcal{B})$$



$$\mathcal{L}(\theta | \gamma, \mathcal{B}) = c \theta^s (1-\theta)^{h-s}$$

$$\ell(\theta | \gamma, \mathcal{B}) = \log c +$$

$$s \log \theta + (h-s) \log (1-\theta)$$

$$\frac{d}{d\theta} \ell(\theta | \gamma, \mathcal{B}) = \frac{s}{\theta} + \frac{h-s}{1-\theta} (-1) = 0$$

$$= \frac{s(1-\theta) - (h-s)\theta}{\theta(1-\theta)} = 0 \text{ iff}$$

$$5 - \cancel{5\theta} - n\theta + \cancel{5\theta} = 0 \quad (6)$$

(5) Fisher (Fred wants constant)

$$\theta = \hat{\theta}_{MLE} = \frac{5}{n} = \bar{y}$$

$$SE_{RS}(\hat{\theta}_{MLE}) = \sqrt{\hat{V}_{RS}(\hat{\theta}_{MLE})}$$

he wants a formula relating

$$\hat{V}_{RS}(\hat{\theta}_{MLE}) \text{ to } \mathcal{L}(\theta | y, R)$$
