

today: poisson;  
bivariate  
uncertainty

C53: Lo5

AMS 206  
19 Feb 19

our uncertainty  
about  
 $\gamma = (\gamma_1, \dots, \gamma_n)$

①

is exchangeable  $\rightarrow$

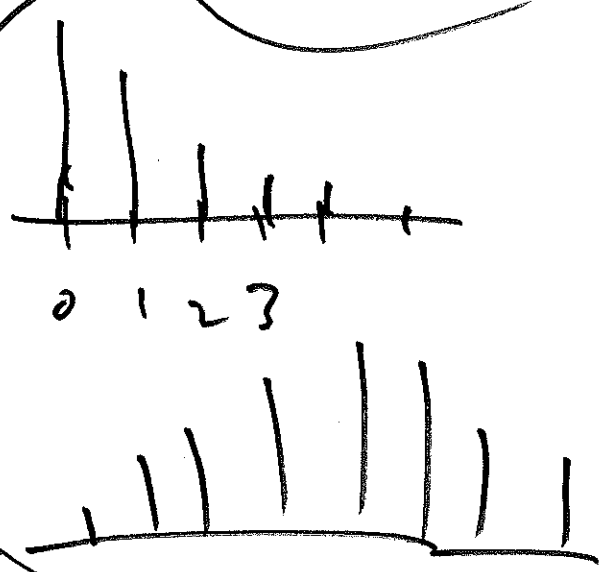
$n=14$

$\left\{ \begin{array}{l} ? \\ (I_i | ?) \sim \text{IID} \\ ? \end{array} \right\}$

$\gamma_i$  non-negative integers

$\swarrow$  poisson  
 $(I_i | \lambda P B) \sim \text{IID}$   
poisson( $\lambda$ )

$(\lambda > 0)$



for integer  $\gamma_i \geq 0$  (marginal)

① (joint)

$$P_{I_i}(\gamma_i | \lambda P B) = \frac{\lambda^{\gamma_i} e^{-\lambda}}{\gamma_i!}$$

$$P_{I_1, \dots, I_n}(\gamma_1, \dots, \gamma_n | \lambda P B) = \prod_{i=1}^n \frac{\lambda^{\gamma_i} e^{-\lambda}}{\gamma_i!}$$

$$= \frac{\lambda^s e^{-n\lambda}}{\prod_{i=1}^n \gamma_i!}$$

$$\textcircled{2} \quad (\lambda > 0)$$

$$\left( s = \sum_{i=1}^n \gamma_i \right) \textcircled{2}$$

$$L(\lambda | Y, P(B))$$

$$= C \lambda^s e^{-n\lambda}$$

$$\textcircled{3} \quad \ell(\lambda | Y, P(B)) = s \log \lambda - n\lambda$$

$(s)$   $(s)$  is sufficient for  $\lambda$  ( $n$ )

$$\textcircled{4} \quad \frac{d}{d\lambda} \ell(\lambda | Y, P(B)) = \frac{s}{\lambda} - n = 0$$

$$\rightarrow \hat{\lambda}_{MLE} = \frac{s}{n} = \bar{Y} = \frac{1}{n} \sum_{i=1}^n \gamma_i = 2.07$$

$$\textcircled{5} \quad \frac{d^2}{d\lambda^2} \ell(\lambda | Y, P(B)) = -\frac{s}{\lambda^2}$$

$$\hat{I}(\hat{\lambda}_{MLE}) = \left[ -\frac{d^2}{d\lambda^2} \ell(\lambda | Y, P(B)) \right]_{\lambda = \hat{\lambda}_{MLE}}$$

$$= \begin{pmatrix} \frac{S}{\lambda^2} \end{pmatrix}$$

$$\lambda = \frac{S}{n}$$

$$= \frac{S}{\frac{S^2}{n^2}} = \frac{n^2}{S}$$

③

$$\vec{V}_{RS}(\hat{\lambda}_{MLE}) = \vec{I}_{RS}^{-1}(\hat{\lambda}_{MLE}) = \frac{1}{n} = O(n^{-1})$$

$$SE_{RS}(\hat{\lambda}_{MLE}) = \sqrt{\vec{V}_{RS}(\hat{\lambda}_{MLE})} = \sqrt{\frac{1}{n}}$$

approx. 95% CI for  $\lambda$

$$\hat{\lambda}_{MLE} \pm 1.96 SE_{RS}(\hat{\lambda}_{MLE})$$

(d = 0.05)  $\vec{I}^{-1}(1 - \frac{\alpha}{2})$

Bayesian algorithm

① ② ③ same as ML story  
 ↓  
 maximum likelihood  
 $L(\lambda | Y, P, B) = c \lambda^S e^{-n\lambda}$

④ is there a conjugate prior? **yes**:  $\Gamma(\alpha, \beta)$  ( $S > 0, n > 0$ )

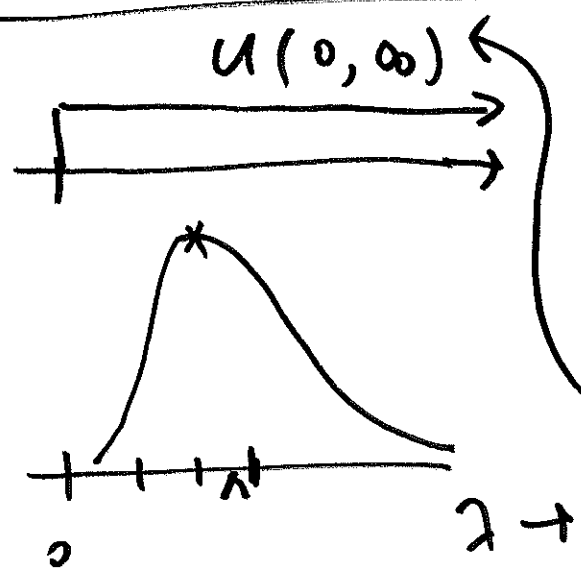
sidebar / (familiarize yourself with

the  $\Gamma(\alpha, \beta)$  family of PDFs

as  $(\alpha > 0, \beta > 0)$  vary.

⑤ conjugate updating

$$c \lambda^{(\alpha+s)-1} e^{-(\beta+y)\lambda} = \underbrace{\left[ c \lambda^{\alpha-1} e^{-\beta\lambda} \right]}_{\text{prior}} \underbrace{\left[ c \lambda^s e^{-y\lambda} \right]}_{\text{likelihood}}$$

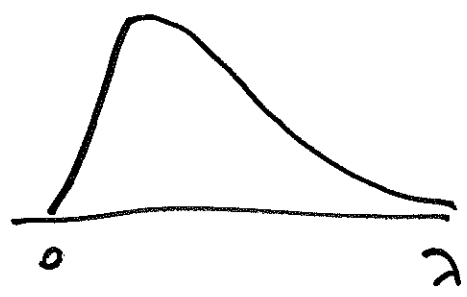


little info about  $\lambda$  external to  $y$

$p(\lambda | y, \alpha, \beta)$

an improper prior for  $\lambda$

$\int_0^{\infty} p(\lambda | \beta) d\lambda = \infty$



$p(\lambda | y, \alpha, \beta)$

$\Gamma(\alpha, \beta)$  prior ; prior sample size  $\beta$

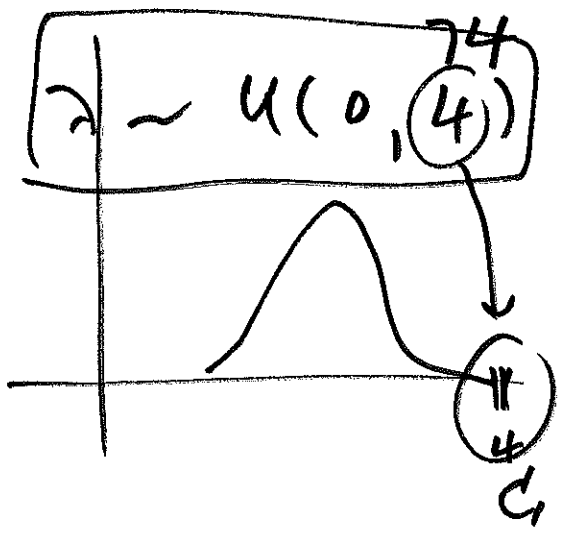
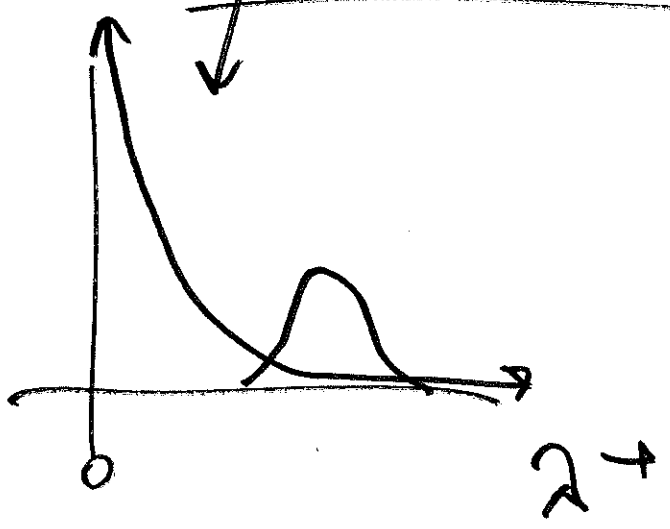
so just choose  $\beta$  small but  $> 0$ ,

e.g.  $\beta = \epsilon = 0.01$

prior  
lots  
set : prior  
mean  
of  $\alpha$   
 $\beta$

common lazy prior (diffuse, flat) but proper  $\rightarrow$  non-informative  
for shy parameter, e.g.  $\lambda > 0$

$\Gamma(\epsilon, \epsilon)$  prior (prior lots set with  $\epsilon$  notes & mean of  $(1)$ )



Description ✓  
 Inference ✓  
 Prediction ✓  
 Decision

$$(\lambda | \alpha, \beta, \Gamma) \sim \Gamma(\alpha, \beta) \textcircled{1}$$

$$(\Gamma_i | \lambda, \mathcal{P}, \mathcal{B}) \stackrel{\text{IID}}{\sim} \text{Poisson}(\lambda)$$

$$(y_1, \dots, y_n) \quad (\Gamma^{-1} = 1, \dots, n)$$

$$(\lambda | y, \mathcal{P}, \Gamma, \mathcal{B}) \sim \Gamma(\alpha + s, \beta + n)$$

posterior predictive distribution for

$$p(y_{n+1} | y, \mathcal{P}, \Gamma, \mathcal{B})$$

a new data value  
 low if total prob.  
 ↓  
 (LTP)

$$= \int_0^{\infty} p(y_{n+1}, \lambda | y, \mathcal{P}, \Gamma, \mathcal{B}) d\lambda$$

$$= \int_0^{\infty} \underbrace{p(y_{n+1} | \lambda, \mathcal{P}, \Gamma, \mathcal{B})}_{\text{canceled}} \boxed{p(\lambda | y, \mathcal{P}, \Gamma, \mathcal{B})} d\lambda$$

$$= \int (\text{Poisson sampling dist}) (\Gamma \text{ posterior})$$